## STAT 542: Statistical Learning

#### Linear Models for Regression

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- Linear Regression Review
- Training vs. Testing Errors
- Model Selection Criteria and Algorithms

### **Linear Models for Regression**

• Observe a collection of n i.i.d. training samples

$$\mathcal{D}_n = \{x_i, y_i\}_{i=1}^n$$

where  $x_i$  is a p dimensional vector (predictors, covariates, features, inputs, i.e.)

$$x_i = (x_{i1}, \dots, x_{ip})^\mathsf{T}$$

and  $y_i \in \mathbb{R}$  is a continuous response (outcome, output).

- We assume the underlying model  $Y = f(X) + \epsilon$
- Estimate f using  $\hat{f}$

•  $\mathbf{x}_j$  is a *n* dimensional vector of the *j*th feature, i.e.

$$\mathbf{x}_j = (x_{1j}, x_{2j}, \dots, x_{nj})^\mathsf{T}$$

• The design matrix  $\mathbf X$  is  $n \times p$ :

$$\mathbf{X} = (\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_p)$$

•  $\mathbf{x}_j$  is one column in  $\mathbf{X}$ 

- What is a good model fitting?
- Loss function, risk, and empirical risk
- A loss function L measures the discrepancy between Y and any function  $f(\boldsymbol{X})$
- · Risk is the expected loss over the entire population

$$\mathsf{R}(f) = \mathsf{E}\left[L(Y, f(X))\right]$$

• The true function f(x) would minimize this risk.

• In regression, the squared error loss is commonly used:

$$L(Y, f(X)) = (Y - f(X))^{2}$$
$$R(f) = E\left[(Y, f(X))^{2}\right]$$

- · Other examples for regression: Huber loss
- For classification: 0/1, logistic, hinge, etc.

#### **The Empirical Risk**

• With the training data  $\mathcal{D}_n$ , estimate f(x) by minimizing the empirical risk

$$\mathsf{R}_n(f) = \frac{1}{n} \sum_{i=1}^n L(y_i - f(x_i))$$

$$\widehat{f}(x) = \underset{f \in \mathcal{F}}{\operatorname{arg\,min}} \quad \mathsf{R}_n(f)$$

where  ${\mathcal F}$  is some space of models

· Using the squared error loss for regression, we have

$$\widehat{f} = \underset{f \in \mathcal{F}}{\operatorname{arg\,min}} \quad \frac{1}{n} \sum_{i=1}^{n} (y_i - f(x_i))^2.$$



Estimating f(X), figure from ESL

• A linear regression model assumes a functional form of f

$$f(X) = X^{\mathsf{T}} \boldsymbol{\beta}$$

- Note: set  $X_1 = 1$  as the intercept term.
- · We express the regression problem in the matrix form

$$\mathbf{y}_{n\times 1} = \mathbf{X}_{n\times p}\boldsymbol{\beta}_{p\times 1} + \mathbf{e}_{n\times 1}$$

• Solve  $\beta$  by minimizing the residual sum of squares (RSS)

$$\mathsf{RSS} = \sum_{i=1}^{n} \left( y_i - x_{i1}\beta_1 - \dots - x_{ip}\beta_p \right)^2$$
$$= \|\mathbf{y} - \mathbf{X}\beta\|^2$$
$$= \left( \mathbf{y} - \mathbf{X}\beta \right)^\mathsf{T} \left( \mathbf{y} - \mathbf{X}\beta \right)$$

• The ordinary least squares estimator (OLS) is

$$\widehat{oldsymbol{eta}} = \operatorname*{arg\,min}_{oldsymbol{eta}} \quad ig(\mathbf{y} - \mathbf{X}oldsymbol{eta}ig)^{\mathsf{T}}ig(\mathbf{y} - \mathbf{X}oldsymbol{eta}ig)$$

#### **Linear Regression**



Solving linear regression, figure from ESL

• To estimate  $\beta$ , we set the derivative equal to 0

$$\frac{\partial \mathsf{RSS}}{\partial \beta} = -2\mathbf{X}^{\top}(\mathbf{y} - \mathbf{X}\beta) = 0$$
$$\implies \mathbf{X}^{\top}\mathbf{y} = \mathbf{X}^{\top}\mathbf{X}\beta$$

which is the normal equation.

- We then have, if  $\mathbf{X}^{\top}\mathbf{X}$  is invertible,

$$\widehat{\boldsymbol{\beta}} = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{y}.$$

-  $\mathbf{X}$  full rank  $\Longleftrightarrow \mathbf{X}^\top \mathbf{X}$  invertible

#### Hat Matrix

• The fitted values (i.e., prediction at observed x<sub>i</sub>'s) are

 $\widehat{\mathbf{y}} = \mathbf{X} (\mathbf{X}^{\top} \mathbf{X})^{-1} \mathbf{X}^{\top} \mathbf{y} \doteq \mathbf{H}_{n \times n} \mathbf{y}$ 

- H ("hat matrix") is a project matrix
  - symmetric:  $\mathbf{H}^{T} = \mathbf{H}$
  - idempotent:  $\mathbf{H}\mathbf{H} = \mathbf{H}$
- The residual  $\mathbf{r}_{n\times 1}=\widehat{\mathbf{e}}=\mathbf{y}-\widehat{\mathbf{y}}=(\mathbf{I}-\mathbf{H})\mathbf{y}$
- · r can be used to estimate the error variance

$$\widehat{\sigma}^2 = \frac{1}{n-p} \sum_{i=1}^n r_i^2 = \frac{\mathsf{RSS}}{n-p}$$

- The essential machinery of linear regression is projection
- Decompose the outcome vector  ${\bf y}$  into two orthogonal vectors

$$\mathbf{y} = \widehat{\mathbf{y}} + \mathbf{r}$$

- $\widehat{\mathbf{y}}$  lives in the column space of  $\mathbf{X}$ , since  $\widehat{\mathbf{y}} = \mathbf{X}\widehat{\boldsymbol{\beta}}$
- **r** is orthogonal to **X**, i.e.,  $\mathbf{X}^{\mathsf{T}}\mathbf{r} = \mathbf{0}$

#### **Vector Space Interpretation**



Figure from Wiki



· We assume that the samples are generated from the model

$$Y = X^{\mathsf{T}}\boldsymbol{\beta} + \boldsymbol{\epsilon},$$

where the errors  $\epsilon_i$  are i.i.d. with  $E(\epsilon_i) = 0$  and  $Var(\epsilon_i) = \sigma^2$ 

- Then  $\widehat{\boldsymbol{\beta}}$  is unbiased:  $\mathsf{E}(\widehat{\boldsymbol{\beta}})=\boldsymbol{\beta}$
- Variance-covariance

$$\operatorname{Var}(\widehat{\boldsymbol{\beta}}) = (\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\sigma^2$$

-  $\sigma^2$  can be estimated using  $\widehat{\sigma}^2$ 



- Among all unbiased linear estimators,  $\widehat{\beta}$  has the smallest variance. (Gauss-Markov Theorem)
- · An unbiased linear estimator is defined as

$$\widehat{oldsymbol{eta}} = \mathbf{A}\mathbf{y}, \ \text{and} \ \mathsf{E}(\widehat{oldsymbol{eta}}) = oldsymbol{eta}$$

- Further assuming  $\epsilon$  is normal,  $\widehat{\boldsymbol{\beta}}$  is also UMVUE
- Question: What if we have a biased estimator? Can we trade a little bias for a large reduction in variance?

## **Training vs. Testing Errors**

- In many applications nowadays, we have many explanatory variables, i.e., *p* is large or even *p* ≫ *n*.
  - There are more than 20,000 human protein-coding genes
  - About 10 million single nucleotide polymorphisms (SNPs)
  - Number of subjects, n, is usually in hundreds or thousands
- In some applications, the key question is to identify a subset of *X* variables that are most relevant to *Y*
- · Let's examine the training and testing errors from a linear model

#### Training vs. Testing error

• Training data  $\mathcal{D}_n = \{x_i, y_i\}_{i=1}^n$ 

- Suppose  $\{x_i, y_i^*\}_{i=1}^n$  is an independent (imaginary) testing dataset collected at the same location  $x_i$ 's (aka, in-sample prediction
- · Assume that the data are indeed from a linear model

$$\mathbf{y} = \boldsymbol{\mu} + \mathbf{e} = \mathbf{X}\boldsymbol{eta} + \mathbf{e}$$
  
 $\mathbf{y}^* = \boldsymbol{\mu} + \mathbf{e}^* = \mathbf{X}\boldsymbol{eta} + \mathbf{e}^*$ 

where both y and y<sup>\*</sup> are  $n \times 1$  response vectors, e and e<sup>\*</sup> are i.i.d. error terms with mean 0 and variance  $\sigma^2$ .

• The true model is indeed linear:  $\mu = X\beta$ 

$$\begin{aligned} \mathsf{E}[\mathsf{Test}\;\mathsf{Err}] &= \mathsf{E}\|\mathbf{y}^* - \mathbf{X}\widehat{\boldsymbol{\beta}}\|^2 \\ &= \mathsf{E}\|(\mathbf{y}^* - \mathbf{X}\boldsymbol{\beta}) + (\mathbf{X}\boldsymbol{\beta} - \mathbf{X}\widehat{\boldsymbol{\beta}})\|^2 \\ &= \mathsf{E}\|\mathbf{y}^* - \boldsymbol{\mu}\|^2 + \mathsf{E}\|\mathbf{X}(\widehat{\boldsymbol{\beta}} - \boldsymbol{\beta})\|^2 \\ &= \mathsf{E}\|\mathbf{e}^*\|^2 + \mathsf{Trace}(\mathbf{X}^\mathsf{T}\mathbf{X}\mathsf{Cov}(\widehat{\boldsymbol{\beta}})) \\ &= n\sigma^2 + p\sigma^2 \end{aligned}$$

$$\begin{split} \mathsf{E}[\mathsf{Train}\;\mathsf{Err}] &= \mathsf{E} \|\mathbf{y} - \widehat{\mathbf{y}}\|^2 = \mathsf{E} \|(\mathbf{I} - \mathbf{H})\mathbf{y}\|^2 \\ &= \mathsf{E} \|(\mathbf{I} - \mathbf{H})\mathbf{e}\|^2 \\ &= \mathsf{Trace}\big((\mathbf{I} - \mathbf{H})^\mathsf{T}(\mathbf{I} - \mathbf{H})\mathsf{Cov}(\mathbf{e})\big) \\ &= (n - p)\sigma^2 \end{split}$$

• Hence, the testing error increases with *p* and training error decreases with *p*. As *p* gets larger, this could be a big trouble...

- It might be necessary to select a set of most relevant variables, especially when *p* is large.
- · Variable selection may improve
  - Prediction accuracy
  - Interpretability
- This is a difficult task
  - No natural ordering of importance for the variables
  - The role of a variable needs be measured conditioning on others, high correlation causes trouble
  - It is essential to check all possible combinations, however, this may be computationally expensive

- · Model selection is usually done in the following way
  - 1) Give each model a score
  - 2) Design an algorithm to find the model with the best (smallest) score
- The score of a model fitting takes the the form

Goodness-of-fit + Complexity-Penalty

- 1) The first term decreases as the model gets more complicated (recall 1NN)
- 2) The second term increases with the number of predictor variables (recall degrees of freedom), which prefers "smaller" model

#### **Model Selection Criteria**

- Popular choices of scores:
  - Mallows'  $C_p$  (Mallows 1973): RSS +  $2\widehat{\sigma}_{\text{full}}^2 \cdot p$
  - AIC (Akaike 1970):  $-2 \text{ Log-likelihood} + 2 \cdot p$
  - BIC (Schwarz, 1978): -2 Log-likelihood  $+ \log n \cdot p$
- When *n* is large, adding an additional predictor costs a lot more in BIC than AIC (or *C<sub>p</sub>*). So AIC tends to pick a larger model than BIC.
- $C_p$  performs similarly to AIC.

- Recall our previous analysis of the training and testing errors with  $\mathbf{y}$  and  $\mathbf{y}^*$
- Now, lets assume that the model is not necessarily a linear model, i.e.,

 $\mathbf{y} = \boldsymbol{\mu} + \mathbf{e}$  $\mathbf{y}^* = \boldsymbol{\mu} + \mathbf{e}^*$ 

We assume mean 0 and variance  $\sigma^2$  for the two error vectors, but we don't have  $\mu = \mathbf{X}\beta$ . However, we still perform linear regression regardless. This will introduce bias of the estimations.

#### Justification of Mallows' $C_p$

$$\begin{aligned} \mathsf{E}[\mathsf{Test}\;\mathsf{Err}] &= \mathsf{E}\|\mathbf{y}^* - \mathbf{X}\widehat{\boldsymbol{\beta}}\|^2 = \|\mathbf{y}^* - \mathbf{H}\mathbf{y}\|^2 \\ &= \mathsf{E}\|(\mathbf{y}^* - \boldsymbol{\mu}) + (\boldsymbol{\mu} - \mathbf{H}\boldsymbol{\mu}) + (\mathbf{H}\boldsymbol{\mu} - \mathbf{H}\mathbf{y})\|^2 \\ &= \mathsf{E}\|\mathbf{y}^* - \boldsymbol{\mu}\|^2 + \mathsf{E}\|\boldsymbol{\mu} - \mathbf{H}\boldsymbol{\mu}\|^2 + \mathsf{E}\|\mathbf{H}\boldsymbol{\mu} - \mathbf{H}\mathbf{y}\|^2 \\ &= \mathsf{E}\|\mathbf{e}^*\|^2 + \mathsf{E}\|\boldsymbol{\mu} - \mathbf{H}\boldsymbol{\mu}\|^2 + \mathsf{E}\|\mathbf{H}\mathbf{e}\|^2 \\ &= \mathbf{n}\sigma^2 + \mathsf{Bias}^2 + p\sigma^2 \end{aligned}$$

$$\begin{aligned} \mathsf{E}[\mathsf{Train}\;\mathsf{Err}] &= \mathsf{E}\|\mathbf{y} - \widehat{\mathbf{y}}\|^2 = \mathsf{E}\|(\mathbf{I} - \mathbf{H})\boldsymbol{\mu} + (\mathbf{I} - \mathbf{H})\mathbf{e}\|^2 \\ &= \mathsf{E}\|(\mathbf{I} - \mathbf{H})\boldsymbol{\mu}\|^2 + \mathsf{E}\|(\mathbf{I} - \mathbf{H})\mathbf{e}\|^2 \\ &= \mathsf{Bias}^2 + (n - p)\sigma^2 \end{aligned}$$

Hence, Test Err is approximately Train Err  $+ 2\sigma^2 p$ , which justifies Mallows'  $C_p$ .

# Model Selection Criteria and Algorithms

- To perform linear model selection, we need to decide on a selection criterion and use an computational algorithm to find the solution
  - Criteria: Mallows' C<sub>p</sub>; AIC; BIC
  - Algorithm: Best subset (Brute force); stepwise (forward/backward/...) selection
- · Different algorithms may have different advantage

- Best subset selection is a level-wise search algorithm, which returns the global optimal solution for a given model size.
- Only feasible for p not very large (< 50)
- Algorithm:
  - 1 For each  $k = 1, \ldots, p$ , check  $2^k$  possible combinations, and find the model with smallest RSS
    - The penalty term is the same for models with the same size
  - 2 To choose the best k, use model selection criteria

- Note: if RSS(X<sub>1</sub>, X<sub>2</sub>) < RSS(X<sub>3</sub>, X<sub>4</sub>, X<sub>5</sub>, X<sub>6</sub>) then we do not need to visit any size 2 or 3 sub-models of (X<sub>3</sub>, X<sub>4</sub>, X<sub>5</sub>, X<sub>6</sub>), which can be leaped over.
- Implemented in R package leaps, using the leaps and bounds algorithm by Furnival and Wilson (1974)

- The Diabetes Data (Efron et al, 2004) contains ten baseline variables from 442 subjects: age, sex, body mass index, average blood pressure, and six blood serum measurements
- The goal is to model a quantitative measure of disease progression one year after baseline
- Data can be loaded from the R package "lars"
- We perform model selections on this dataset (see R code from course material)

- Greedy algorithms: fast, but only return a local optimal solution (which might be good enough in practice).
  - Backward: start with the full model and sequentially delete predictors until the score does not improve.
  - Forward: start with the null model and sequentially add predictors until the score does not improve.
  - Stepwise: consider both deleting and adding one predictor at each stage.